

# Plasma sex hormone concentrations and breast cancer risk in an ethnically diverse population of postmenopausal women: the Multiethnic Cohort Study

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## Abstract

To add to the existing evidence that comes mostly from White populations, we conducted a nested case–control study to examine the association between sex hormones and breast cancer risk within the Multiethnic Cohort that includes Japanese American, White, Native Hawaiian, African American, and Latina women. Of the postmenopausal women for whom we had a plasma sample, 132 developed breast cancer during follow-up. Two controls per case, matched on study area (Hawaii, Los Angeles), ethnicity/race, birth year, date and time of blood draw and time fasting, were randomly selected from the women who had not developed breast cancer. Levels of estradiol (E<sub>2</sub>), estrone (E<sub>1</sub>), androstenedione, dehydroepiandrosterone (DHEA), and testosterone were quantified by RIA after organic extraction and Celite column partition chromatography. E<sub>1</sub> sulfate, DHEA sulfate (DHEAS), and sex hormone-binding globulin (SHBG) were quantified by direct immunoassays. Based on conditional logistic regression, the sex hormones were positively associated and SHBG was negatively associated with breast cancer risk. All associations, except those with DHEAS and testosterone showed a significant linear trend. The odds ratio (OR) associated with a doubling of E<sub>2</sub> levels was 2.26 (95% confidence interval (CI) 1.58–3.25), and the OR associated with a doubling of testosterone levels was 1.34 (95% CI 0.98–1.82). The associations in Japanese American women, who constituted 54% of our sample, were similar to or nonsignificantly stronger than in the overall group. This study provides the best evidence to date that the association between sex hormones and breast cancer risk is generalizable to an ethnically diverse population.

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## Introduction

Many risk factors for postmenopausal breast cancer are thought to be mediated by a hormonal mechanism (Henderson & Feigelson 2000, Clemons & Goss 2001). The investigation of circulating sex hormones has provided evidence that is more direct. A pooled analysis published in 2002 (Key *et al.* 2002) of nine prospective studies of women not using hormone replacement therapy (HRT) found that all estrogenic and androgenic steroid sex hormones examined were associated with increased breast cancer risk, and

that sex hormone-binding globulin (SHBG) was associated with decreased risk. Since the publication of this pooled analysis, these findings have been supported for the most part by analyses in some of the studies with extended follow-up (Missmer *et al.* 2004, Zeleniuch-Jacquotte *et al.* 2004, Cummings *et al.* 2005, Eliassen *et al.* 2006, Sieri *et al.* 2009) and additional large prospective studies (Cummings *et al.* 2002, Manjer *et al.* 2003, Onland-Moret *et al.* 2003, Kaaks *et al.* 2005, Beattie *et al.* 2006, Gunter *et al.* 2009).

With a single exception, all of these studies were conducted in largely White populations. Kabuto *et al.* (2000) investigated the association of breast cancer with estradiol ( $E_2$ ), SHBG, and dehydroepiandrosterone sulfate (DHEAS) in a prospective study of Japanese women in Japan, but had only 26 postmenopausal cases and did not investigate testosterone levels. The odds ratio (OR) for the association with a doubling of  $E_2$  levels was 0.90 (95% confidence interval (CI) 0.42–1.92) (reported in Key *et al.* (2002)) suggesting that the association may be different in women of Japanese ancestry. Two case–control studies conducted in China, on the other hand, found that breast cancer risk in non-Japanese Asian postmenopausal women was positively associated with estrogen and testosterone levels (Yu *et al.* 2003, Wang *et al.* 2009). This report examines the association between estrogens ( $E_2$ , estrone ( $E_1$ ), and  $E_1$  sulfate), androgens (androstenedione, DHEA, DHEAS, and testosterone), and SHBG and breast cancer among postmenopausal women not using HRT within the Multiethnic Cohort (MEC) study. Because so few results are available from past studies in Asian populations, another objective was to examine these associations in the subgroup restricted to our Japanese American participants who comprise a large proportion of the MEC.

## Materials and methods

### Study design and population

We conducted a case–control study nested within the MEC. This cohort includes over 215 000 men and women between the ages of 45 and 75 years at recruitment living in Hawaii and Los Angeles, California who returned a baseline questionnaire between 1993 and 1996 (Kolonel *et al.* 2000). From information in the questionnaire, participants were assigned to a single ethnic/racial group and if more than one group was reported, assignment was based on the following priority ranking: African American, Native Hawaiian, Latino, Japanese American, and White. More than 95% of all non-Hawaiians reported only one ethnic group. Body mass index (BMI) was calculated as weight divided by the square of height ( $\text{kg}/\text{m}^2$ ). Women were defined as postmenopausal if they reported cessation of menstrual periods naturally or due to bilateral oophorectomy, reported a hysterectomy without bilateral oophorectomy and use of HRT, or if this information was missing, were aged 55 years or older at blood draw.

Largely between 2001 and 2006, 67 594 MEC participants agreed to donate a biospecimen and

complete a short questionnaire that updated information including HRT use and menopausal status. Fasting blood was drawn in the early morning at a clinical laboratory or in the participant's home, processed, and separated into components. The plasma for the present hormone assays was stored in 2 ml aliquots at  $-80^\circ\text{C}$ . Some women only had serum available in our repository and were excluded because the volume of serum that could be extracted from the 0.5 ml cryotubes in which it was stored was too low to reliably quantify the low hormone levels found in postmenopausal women. The MEC and this study were approved by the Institutional Review Boards of the University of Southern California and the University of Hawaii. All participants provided informed consent at the time of blood collection.

### Cases and controls

Cases were women who had contributed a blood specimen before being diagnosed with a first primary breast cancer to the end of follow-up (December, 2006) and had plasma available for analysis. Diagnoses were identified by linkage of the MEC to the National Cancer Institute's Surveillance, Epidemiology, and End Results (SEER) program that covers the entire population of the Hawaii and California. Deaths were identified by linkage with state vital statistics databases and the National Death Index. Two control women per case were randomly selected from a pool of subcohort members who were alive and not diagnosed with breast cancer at the same age as the case at diagnosis, and matched to the case by area (Hawaii, Los Angeles), ethnicity, birth year ( $\pm 1$  year), date of blood draw ( $\pm 6$  months), time of blood draw ( $\pm 2$  h), hours fasting before blood draw (0–<6, 6–<8, 8–<10,  $\geq 10$ ), and HRT use at the time of blood draw. Only women not using HRT at the time of blood collection were included in this analysis.

Of the 326 cases who had a fasting blood sample available in the biorepository, 124 were currently using HRT and 70 had serum samples only. Thus, 132 incident breast cancer cases not using HRT had a plasma sample available for hormone measurement. Of the 264 controls matched to these cases, one did not have a plasma sample and two did not have sufficient sample for assays. Thus, a total of 132 cases and 261 controls had assay results for the analytes of primary interest (estrogens, testosterone, and DHEA). One control did not have sufficient sample to assay  $E_1$  sulfate and DHEAS.

## Laboratory assays

Assays for sex hormones and SHBG were performed on plasma samples at the Reproductive Endocrine Research Laboratory at the University of Southern California (FZS). Each batch included matched case–control sets and pooled quality control samples, which were randomly ordered and were blinded to laboratory personnel, plus laboratory standards. Levels of the unconjugated steroid hormones were determined by RIA after organic extraction and partition chromatography on Celite columns (Goebelsmann *et al.* 1973, Stanczyk *et al.* 2007, 2009). Levels of E<sub>1</sub> sulfate were determined by direct immunoassay using kits from Beckman-Coulter Diagnostic Systems Laboratories (Webster, TX, USA; Ranadive *et al.* 1998). Levels of DHEAS and SHBG were determined using a chemiluminescent immunoassay on an Immulite analyzer (Siemens Medical Solutions USA Inc., Malvern, PA, USA). The limits of detection were 5 pg/ml for E<sub>1</sub>, 0.01 ng/ml for E<sub>1</sub> sulfate, 2 pg/ml for E<sub>2</sub>, 0.03 ng/ml

for androstenedione, 0.04 ng/ml for DHEA, 15 µg/dl for DHEAS, 1.5 ng/dl for testosterone, and 1 nmol/l for SHBG. Seven women were below the limit for E<sub>2</sub> and 57 women were below the limit for DHEAS. These women were assumed to have levels equal to half of the detection limit. Bioavailable testosterone and E<sub>2</sub> were estimated with a validated algorithm using the measured concentrations of the sex hormones and SHBG and an assumed albumin concentration of 43 g/l (Sodergard *et al.* 1982, Vermeulen *et al.* 1999, Rinaldi *et al.* 2002). The intra-batch coefficients of variation (CV), based on 15 paired duplicates, were 10% for E<sub>1</sub>, 8.5% for E<sub>1</sub> sulfate, 8.2% for E<sub>2</sub>, 5.9% for androstenedione, 3.9% for DHEA, 4.6% for DHEAS, 4.2% for testosterone, and 7.5% for SHBG.

## Statistical analysis

ORs for the association between breast cancer risk and hormone concentration in quartiles based on the distribution in controls were computed using

**Table 1** Characteristics of the cases and controls

Characteristic	All women		Japanese American	
	Cases N (%)	Controls N (%)	Cases N (%)	Controls N (%)
Number of subjects	132 (100)	261 (100)	72 (100)	142 (100)
Ethnicity/race <sup>a</sup>				
Japanese American	72 (54.5)	142 (54.4)	72 (100)	142 (100)
White	30 (22.7)	60 (23.0)	0 (0)	0 (0)
Native Hawaiian	17 (12.9)	34 (13.0)	0 (0)	0 (0)
African American	8 (6.1)	15 (5.7)	0 (0)	0 (0)
Latina	5 (3.8)	10 (3.8)	0 (0)	0 (0)
Number of children <sup>b</sup>				
0	17 (12.9)	28 (10.8)	6 (8.3)	11 (7.7)
1	7 (5.3)	24 (9.2)	4 (5.6)	17 (12.0)
2–3	70 (53.0)	134 (51.5)	45 (62.5)	86 (60.6)
≥4	38 (28.8)	74 (28.5)	17 (23.6)	27 (19.0)
Aged ≥25 years at first birth, parous women <sup>b</sup>	40 (34.8)	65 (28.0)	28 (42.4)	50 (38.5)
Residence in Hawaii <sup>a</sup>	113 (85.6)	224 (85.8)	66 (91.7)	130 (91.5)
> 12 years of education <sup>b</sup>	91 (68.5)	172 (65.9)	52 (72.2)	101 (71.1)
Menarche at <13 years of age <sup>b</sup>	70 (53.8)	145 (56.4)	40 (55.6)	79 (55.6)
Past oral contraceptive use	67 (50.8)	119 (45.6)	32 (44.4)	60 (42.3)
Past use of HRT <sup>b</sup>	68 (52.7)	132 (51.2)	37 (51.4)	71 (50.0)
First degree family history of breast cancer	10 (7.6)	34 (13.0)	5 (6.9)	16 (11.3)
Consumed ≥1 alcoholic drink/day at baseline	12 (9.1)	23 (8.8)	1 (1.4)	3 (2.1)
	<b>Median (IQR)</b>	<b>Median (IQR)</b>	<b>Median (IQR)</b>	<b>Median (IQR)</b>
Age at blood draw <sup>a</sup> (years)	68.0 (61.4–75.7)	68.1 (61.2–75.2)	67.7 (62.7–74.0)	67.6 (62.0–73.4)
Years between blood draw and case diagnosis <sup>a</sup>	0.9 (0.5–1.9)	1.0 (0.5–1.9)	1.3 (0.6–2.3)	1.3 (0.6–2.3)
Years between baseline and blood draw	9.2 (8.6–10.4)	9.4 (8.9–10.4)	9.0 (8.5–10.4)	9.3 (8.8–10.1)
Hours fasting before blood draw <sup>a</sup>	13.1 (12.0–14.1)	12.9 (11.8–14.1)	13.3 (12.0–14.1)	13.1 (11.8–14.0)
Body mass index at baseline <sup>b</sup> (kg/m <sup>2</sup> )	24.7 (22.4–27.6)	24.0 (21.5–27.5)	24.0 (21.5–26.7)	23.2 (20.9–25.8)
Physical activity at baseline <sup>b</sup> (MET-hours/week)	4.2 (1.4–8.9)	3.6 (2.0–8.2)	4.8 (1.4–8.9)	3.6 (2.0–8.0)

HRT, hormone replacement therapy; IQR, interquartile range (25th to 75th percentile); MET, metabolic equivalent.

<sup>a</sup>Matching variables.

<sup>b</sup>Summary statistics based on slightly fewer than the total N in each group due to missing values.

conditional logistic regression with the matched sets as strata. Because levels of hormones and SHBG were positively skewed, they were logarithmically (base 2) transformed when used as continuous variables as was done in the pooled analysis (Key *et al.* 2002). The corresponding ORs represent the risk associated with a doubling of hormone concentration; these analyses also constituted our test for trend. Included in the models were two of the matching factors as continuous variables, age at blood draw and hours fasting, to control for any residual differences in these variables between cases and matched controls. We also evaluated a number of risk factors as potential confounders (listed in Table 1). None was retained as a confounder because when added to the models, they changed the ORs associated with a doubling of hormone concentration no more than 10% (Mickey & Greenland 1989). Analyses stratified by HRT use (past, never) and time between blood collection and diagnosis of the case in the set ( $<1$ ,  $\geq 1$  year) were done. These analyses used hormone levels as continuous variables; the product between the stratification variable and hormone variables was introduced into the models to assess the significance of any effect modification. An analysis stratified by ethnicity, using hormone levels as continuous variables, compared Japanese American women to the group of women of other ethnicities/races combined. We did not have a sufficient number of women in any other single ethnic/racial group to compare directly to the group of Japanese American women. The analyses were done using SAS version 9.1 (SAS Institute, Cary, NC, USA). All statistical tests were two-sided with a 0.05 level of significance.

## Results

Characteristics of all women (132 cases and 261 controls) and the subgroup of Japanese American

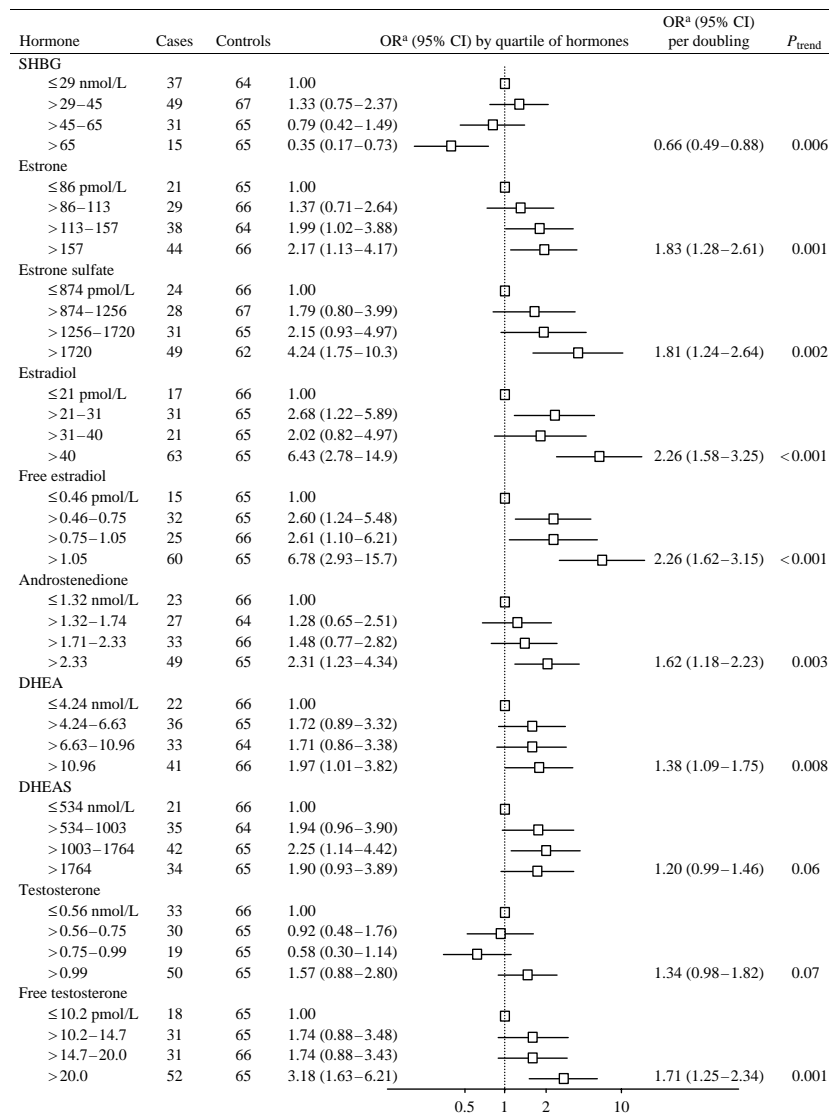
women (72 cases and 142 controls) are shown in Table 1. The median age of all participants was 68 years, and over half of the women were Japanese American. About half of both cases and controls used HRT in the past. Levels of the sex hormones were uniformly higher in cases than controls, whereas the level of SHBG was lower in cases than controls (Table 2). Among controls, levels of the analytes were often significantly correlated. For example, the Pearson correlations (using log-transformed values) of  $E_2$  with  $E_1$ , testosterone, and SHBG were 0.51, 0.33, and  $-0.27$  respectively (all  $P < 0.05$ ). Among controls, the geometric mean  $E_2$  concentration adjusted for age and BMI varied significantly among the ethnic/racial groups ( $P = 0.03$ ), with African American women having high average levels (39 pmol/l) and Latinas having low average levels (20 pmol/l) relative to the other three groups (28 pmol/l).

Sex hormones were positively associated with breast cancer risk, and SHBG was inversely associated with breast cancer risk (Fig. 1). All associations, except those with DHEAS and testosterone, showed a significant trend, the strongest of which were observed for  $E_2$  and free  $E_2$ . Women who had  $E_2$  levels in the highest quartile had over sixfold increased risk of breast cancer relative to women in the lowest quartile (OR 6.43, 95% CI 2.78–14.9). We did not find that past HRT use significantly modified any of these associations (all  $P$  for heterogeneity  $> 0.05$ ). Generally, the trends observed in the overall group were reflected in both past and never users of HRT but were larger in magnitude among the former group (data not shown). The ORs for the associations between a doubling of estrogen levels and breast cancer risk were slightly, but nonsignificantly, greater if blood was taken 1 year or more before diagnosis whereas the OR for the association between SHBG and breast cancer risk

**Table 2** Geometric mean (95% confidence interval) concentrations of sex hormones and SHBG in cases and controls

Sex hormone/SHBG	All women		Japanese American	
	Cases	Controls	Cases	Controls
SHBG (nmol/l)	38 (35–41)	44 (41–47)	35 (31–40)	43 (39–48)
Estrone (pmol/l)	132 (122–143)	112 (106–118)	128 (115–143)	108 (100–116)
Estrone sulfate (pmol/l)	1363 (1218–1525)	1168 (1092–1250)	1548 (1372–1748)	1295 (1189–1410)
Estradiol (pmol/l)	37 (33–41)	28 (26–30)	36 (32–41)	28 (25–30)
Free estradiol (pmol/l)	0.96 (0.87–1.07)	0.70 (0.64–0.76)	0.96 (0.83–1.10)	0.68 (0.61–0.75)
Androstenedione (nmol/l)	2.02 (1.87–2.19)	1.73 (1.63–1.84)	1.99 (1.79–2.21)	1.66 (1.54–1.79)
DHEA (nmol/l)	7.61 (6.82–8.49)	6.30 (5.79–6.86)	8.20 (7.15–9.41)	6.72 (6.07–7.43)
DHEAS (nmol/l)	1068 (927–1229)	916 (825–1018)	1217 (1024–1447)	1071 (931–1231)
Testosterone (nmol/l)	0.83 (0.76–0.90)	0.75 (0.71–0.80)	0.77 (0.70–0.86)	0.71 (0.66–0.77)
Free testosterone (pmol/l)	17.2 (15.9–18.6)	14.4 (13.5–15.3)	16.6 (15.1–18.2)	13.6 (12.42–14.8)

DHEA, dehydroepiandrosterone; DHEAS, dehydroepiandrosterone sulfate; SHBG, sex hormone-binding globulin.



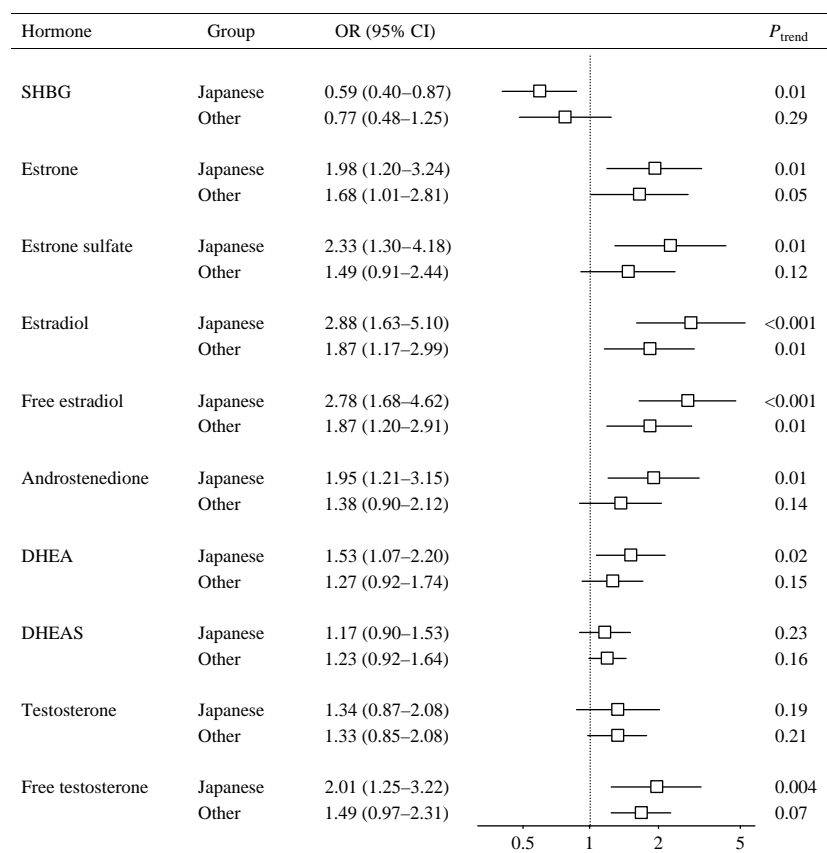
**Figure 1** Odds ratios (OR) and 95% confidence intervals (CI) for the association of hormone concentrations with the risk of breast cancer. <sup>a</sup>Cases and controls were matched on area, ethnicity, birth year ( $\pm 1$  year), date of blood draw ( $\pm 6$  months), time of blood draw and hours fasting (in categories), and ORs were additionally adjusted for age at blood draw and hours fasting as continuous variables.

was slightly, but nonsignificantly, more protective if blood was taken  $< 1$  year before diagnosis (data not shown). As depicted in Fig. 2, when the analysis was restricted to Japanese American women, all associations except those with DHEAS and testosterone showed a significant trend and were nonsignificantly stronger than the associations in women of other ethnicity/race (all  $P$  for heterogeneity  $> 0.05$ ). For example, the OR for a doubling of  $E_2$  was 2.88 (95% CI 1.63–5.10) in Japanese American women as compared to 1.87 (95% CI 1.17–2.99) in women of other ethnicity/race.

## Discussion

In this study, estrogenic and androgenic hormones were positively associated, and SHBG was inversely associated with breast cancer risk. The pooled analysis done in 2002 (Key *et al.* 2002) and four prospective studies of circulating hormones done since then (Cummings *et al.* 2002, Manjer *et al.* 2003, Kaaks *et al.* 2005, Gunter *et al.* 2009) have also found such associations although a fifth prospective study did not (Beattie *et al.* 2006). In the current study, stronger associations than in other prospective studies were observed for some analytes such as  $E_2$ ,  $E_1$  sulfate,





**Figure 2** Odds ratios (OR) and 95% confidence intervals (CI) for the association between a doubling of hormone concentrations and the risk of breast cancer among Japanese American women and women of other race/ethnicity. <sup>a</sup>Cases and controls were matched on area, ethnicity, birth year ( $\pm 1$  year), date of blood draw ( $\pm 6$  months), time of blood draw and hours fasting (in categories), and ORs were additionally adjusted for age at blood draw and hours fasting as continuous variables.

and SHBG. For example, the OR for a doubling of  $E_2$  concentration was 2.26 (95% CI 1.58–3.25), whereas in the pooled analysis, it was 1.29 (95% CI 1.15–1.44) with ORs ranging from 0.76 to 1.69 in the individual studies.

The impact of the short time between blood collection and diagnosis (median of 1 year), which could be construed as a limitation of our study, should be considered. Because mammary tumors express aromatase, they have the capability of producing estrogens from androgens (Pasqualini *et al.* 1996), and if the estrogens thus produced enter the circulation, reverse causality could be a concern. However, in breast cancer patients, breast tissue estrogens do not reflect circulating estrogens (Geisler 2003). Furthermore, in one study that collected two blood samples a mean of 31 months apart, the rate of change in most hormone levels was not significantly different between cases and controls, suggesting that the presence of an undetected tumor does not have a large effect on circulating hormone levels (Zeleniuch-Jacquotte *et al.* 2004).

The second sample, however, was collected on average 28 months before the diagnosis of the case, and the possibility still remains that tumors closer than this to diagnosis have an effect on circulating hormone levels. Additional information about the impact of the time between blood collection and diagnosis comes from prospective studies that have stratified on this factor, which had little effect on most associations (Dorgan *et al.* 1996, Zeleniuch-Jacquotte *et al.* 1997, Hankinson *et al.* 1998, Kabuto *et al.* 2000, Key *et al.* 2002, Sieri *et al.* 2009). Associations between a few of the hormones and breast cancer risk, however, were weaker (Toniolo *et al.* 1995, Dorgan *et al.* 1997, Hankinson *et al.* 1998, Onland-Moret *et al.* 2003) or stronger (Dorgan *et al.* 1996, Kabuto *et al.* 2000, Zeleniuch-Jacquotte *et al.* 2004) with less time between blood collection and diagnosis.

The strength of the association observed for circulating estrogen concentrations with breast cancer risk was notable in our multiethnic population. Some of these associations may have been stronger than in

other prospective studies because our indirect assays have been shown to have high validity and reliability (Lee *et al.* 2006), perhaps leading to less attenuation of risk estimates due to nondifferential exposure misclassification. E<sub>2</sub> measurements done in the laboratory in which our assays were performed correlate very well with the gold standard method gas chromatography tandem mass spectrometry (GC-MS/MS), whereas direct assays without an extraction step that reduces cross-reactivity with other hormone metabolites correlate less well (Lee *et al.* 2006). The CV of the assays done in this study, however, are not noticeably lower than those reported in other studies where case–control sets were placed in the same batch to eliminate the effect of inter-batch variability (Tworoger & Hankinson 2006). Also possibly contributing to the strength of our associations is the fact that Japanese American women comprised over half of our sample in whom the associations between most sex hormones and breast cancer risk appeared to be of greater magnitude than in women of other ethnicity/race. This difference, however, would not be enough to wholly explain our stronger results for E<sub>1</sub>, E<sub>2</sub>, and SHBG in comparison to previous prospective studies.

The only other prospective study done among postmenopausal Asian women, conducted in Japan, found no significant association between breast cancer risk and total E<sub>2</sub>, bioavailable E<sub>2</sub>, DHEAS, and SHBG; the point estimate for E<sub>2</sub> was below 1.0 but the 95% CIs were wide due to a small sample size (Kabuto *et al.* 2000). Two case–control studies conducted in China, however, found positive associations between postmenopausal breast cancer risk and levels of estrogens and testosterone (Yu *et al.* 2003, Wang *et al.* 2009). The Japanese American women in the current study may have had higher levels of sex hormones than the Asian women in the other studies: higher sex hormone levels have been observed in Western Japanese as compared to Eastern Japanese populations in studies using the same assay techniques (Shimizu *et al.* 1990). Although this appears to be contradicted by the reported median or geometric mean E<sub>2</sub> concentration being higher in the controls of two of these studies (Kabuto *et al.* 2000, Wang *et al.* 2009) than in the current study, comparing levels across studies is not possible when assay techniques are different (Key *et al.* 2002).

In the MEC, Japanese American women are at similar or slightly higher risk of breast cancer relative to White women (Pike *et al.* 2002, Woolcott *et al.* 2009). This result is compatible with observations made in the current study population and another within the MEC (Setiawan *et al.* 2006) that Japanese

American women have sex hormone levels similar to or slightly higher than White women once age, BMI, and other factors are taken into account. Postmenopausal Japanese American women in the MEC, of whom 89% were born in the United States, may lead a fairly westernized lifestyle and thus may have a distribution similar to White women of lifestyle factors that affect sex hormone concentration. For example, indicators of a westernized diet, such as low vegetable and soy intake, are associated with lower SHBG concentrations (Wu *et al.* 2009) and higher testosterone concentrations (Setiawan *et al.* 2006). With respect to dietary patterns, however, Japanese Americans in the MEC are still likely to differ somewhat from Whites (Park *et al.* 2005).

In this study, we had sufficient statistical power with the whole sample to detect associations of the magnitude observed in other prospective studies of circulating hormones and breast cancer risk but our ability to detect differences in effect among subgroups was limited by our sample size. We had a larger number of women of Asian ancestry than in other published prospective studies, and we found significant associations in this group. Because the other ethnic/racial groups each had few subjects, we compared the results in the Japanese American women to those in the other groups combined. As follow-up continues, we will be able to pursue analyses by each ethnic/racial group separately. Although we relied on a single blood draw for our analysis, other studies have observed that hormone levels up to 5 years apart are moderately correlated suggesting that a single hormone measurement adequately characterizes levels over a longer period of time (Toniolo *et al.* 1995, Kim & Zeleniuch-Jacquotte 1997, Hankinson *et al.* 1998, Zeleniuch-Jacquotte *et al.* 2004). Our covariate information was largely derived from the baseline questionnaire, which was done an average of 9.3 years before the blood was drawn. This limitation only applies to time-dependent covariates, such as BMI, which are unlikely to change the conclusions in this study because in most other studies, covariates confound the associations between circulating sex hormones and breast cancer risk very little (Dorgan *et al.* 1996, Hankinson *et al.* 1998, Kabuto *et al.* 2000, Key *et al.* 2002, Endogenous Hormones and Breast Cancer Collaborative Group 2003, Onland-Moret *et al.* 2003, Zeleniuch-Jacquotte *et al.* 2004, Cummings *et al.* 2005, Kaaks *et al.* 2005).

In conclusion, this study supports previous evidence that sex hormone levels are associated with an increased risk of breast cancer, and SHBG levels are associated with a decreased risk among

postmenopausal women. It demonstrates that these associations are generalizable to an ethnically diverse population, and furthermore, it suggests that these associations remain similar when restricted to Japanese American women.

### Declaration of interest

The authors declare that there is no conflict of interest that could be perceived as prejudicing the impartiality of the research reported.

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### Author contribution statement

C G Woolcott carried out the statistical analysis, interpreted the results, and wrote the manuscript. L R Wilkens and M T Goodman also helped to interpret the results. Y B Shvetsov and L R Wilkens provided statistical expertise. K K White and C Caberto were involved in data acquisition, management, and subject selection. B E Henderson and L N Kolonel are the Principal Investigators of the Multiethnic Cohort Study and with M T Goodman, L R Wilkens, and L Le Marchand were involved in the concept and procurement of funding for this particular project. F Z Stanczyk carried out the hormone assays and helped to interpret this data. All authors provided feedback on the initial draft of the manuscript and approved the final manuscript.

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